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On Universality of Classical Probability with Contextually Labeled Random Variables

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One can often encounter claims that classical (Kolmogorovian) probability theory cannot handle, or even is contradicted by, certain empirical findings or substantive theories. This note joins several previous attempts to explain that these claims are unjustified, illustrating this on the issues of (non)existence of joint distributions, probabilities of ordered events, and additivity of probabilities. The specific focus of this note is on showing that the mistakes underlying these claims can be precluded by labeling all random variables involved contextually. Moreover, contextual labeling also enables a valuable additional way of analyzing probabilistic aspects of empirical situations: determining whether the random variables involved form a contextual system, in the sense generalized from quantum mechanics. Thus, to the extent the Wang-Busemeyer QQ equality for the question order effect holds, the system describing them is noncontextual. The double-slit experiment and its behavioral analogues also turn out to form a noncontextual system, having the same probabilistic format (cyclic system of rank 4) as the one describing spins of two entangled electrons.

KEYWORDS: classical probability, contextuality, contextual labeling, double-slit experiment, question-order effects, random variables.

In the literature on foundations of quantum physics (Accardi, 1982; Feynman, 1951; Feynman, Leighton, & Sands, 1975; Khrennikov, 2009b) and, more recently, psychology (Aerts, 2009, 2014; Broekaert, Basieva, Blasiak, & Pothos, 2017; Busemeyer & Bruza, 2012; Moreira & Wichert, 2016; Pothos & Busemeyer, 2013), one can encounter statements that classical (Kolmogorovian) probability theory does not have adequate conceptual means to handle (sometimes, even, is contradicted by) this or that empirical fact.

Three of the most widespread assertions of this kind are as follows:

Statement 1: Classical probability requires that certain (e.g., Bell-type) inequalities hold for certain systems of random variables, but we know from quantum mechanics and from behavioral experiments that they may be violated.

Statement 2: In classical probability, the joint occurrence of two events is commutative, but we know from quantum mechanics and from behavioral experiments that the order of two events generally matters for their joint probability.

Statement 3: Classical probability is additive (equivalently, obeys the law of total probability), but we know from quantum mechanics and from behavioral experiments that this additivity (the law of total probability) can be violated.

This note has three objectives: (1) to show that the three statements above are based on misidentification of the random variables involved, due to ignoring their inherently contextual labeling; (2) to show that contextual labeling is a principled way to "automatically" ensure correct applicability of classical probability theory to an empirical

situation; and (3) to demonstrate how the use of contextual labeling enables so-called contextuality analysis of systems of random variables, a relatively new form of probabilistic analysis of considerable interest in empirical applications. Contextual labeling of random variables is the departing principle of Khrennikov's Växjö Model (Khrennikov, 2009a) and of the Contextuality-by-Default theory (Dzhafarov, 2017; Dzhafarov, Cervantes, & Kujala, 2017; Dzhafarov & Kujala, 2014a, 2015, 2016a, 2017a, 2017b; Dzhafarov, Kujala, & Cervantes, 2016).

Let us preamble this discussion by stating our view of classical probability theory (CPT), one that we are not prepared to defend in complete generality, confining ourselves instead to merely illustrating it on the three statements above. This view is that CPT, on a par with classical logic and set theory, is a universal abstract mathematical theory. As an abstract mathematical theory, it does not make empirically testable predictions, because of which it cannot be contradicted by any empirical observation. As a universal theory, for any empirical situation, it has conceptual means to adequately describe anything that can be qualified as this situation's probabilistic features (in the frequentist sense). Moreover, as a conceptual tool, in the same way as classical logic and set theory, it is indispensable and irreplaceable in dealing with probabilistic problems: at the end, the results of any non-classical probabilistic analysis have to be formulated in terms of classical (frequentist) probabilities, distributions, and random variables. However, when applied to an empirical situation, CPT can (even must) be complemented by special-purpose computations identifying some of the random variables, distributions, and probabilities in this particular situation. To give a very simple example, CPT provides methods for deriving probabilities of events defined on the outcomes of rolling a die from a distribution of these outcomes, but it cannot predict this distribution. A special theory is needed to know, e.g., that if a die is manufactured in a particular way, then the distribution of its outcomes is uniform. We view quantum probability as such a special-purpose theory complementing classical probability. This mathematical formalism is indispensable

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in quantum mechanics and has significant achievements to its credit in psychology (e.g., Wang & Busemeyer, 2013). It can be formalized and presented as an abstract calculus alternative to or even generalizing the calculus of CPT, in the same way one can formalize a paraconsistent logic as a generalization of classical logic. However, just as one cannot replace classical logic with paraconsistent logic in analyzing anything, including the very paraconsistent logic itself, one cannot dispense with classical probability when discussing and analyzing quantum probability computations and relating them to data.

This view is not entirely new. Ballentine (1986) defended a similar position in essentially the same way we are doing here. The difference is in that instead of using random variables, Ballentine confined himself to a more limited language of events, and he used conditionalization in place of the more general contextualization (Dzhafarov & Kujala, 2014b; we discuss conditionalization in Section 3 below). Khrennikov (2009a), in describing his Växjö contextual model uses Ballentine's conditional-probability notation, but emphasizes that these are not conditional probabilities of CPT. Rather he calls them "contextual probabilities," and explains that "contextual probability [...] is not probability that an event, say B, occurs under the condition that another event, say C, has occurred. The contextual probability is probability to get the result $a = \alpha$ under the complex of physical conditions C" (Khrennikov, 2009a, p. 50). This seems to be the same as the contextual labeling used in the Contextuality-by-Default theory. A very clear presentation of a position that is close to ours can be found in the arguments presented in an internet discussion by Tim Maudlin (2013).

The purpose of this paper is to achieve conceptual clarity in understanding CPT, not to criticize specific authors or papers. The latter is an ungrateful task, as most authors' positions are not entirely consistent, are subject to (re)interpretations, and evolve over time. We cite specific papers and occasionally provide quotes only to demonstrate that a reasonable reader may interpret the positions they entail in the spirit of the Statements 1-3 above. Thus, Richard Feynman is often cited as arguing that classical probability is not compatible with quantum mechanics (Accardi, 1982; Costantini, 1993; Khrennikov 2009b). This interpretation is supported by Feynman's speaking of "the discovery that in nature the laws of combining probabilities were not those of the classical probability theory of Laplace" (Feynman, 1951, p. 533). However, one can also find statements in Feynman's writings that make his point of view less than unequivocal. Thus, we read in the same paper and on the same page that "the concept of probability is not altered in quantum mechanics. When I say the probability of a certain outcome of an experiment is p[...] no departure from the concept used in classical statistics is required. What is changed, and changed radically, is the method of calculating probabilities" (ibid). This quote is consistent with treating quantum formalisms as specialpurpose computations embedded in CPT. We will return to Feynman when discussing the double-slit experiment in Section 3.

1. ON STATEMENT 1

"Classical probability requires that certain (e.g., Bell-type) inequalities hold for certain sets of random variables, but we know from quantum mechanics and from behavioral experiments that they may be violated."

This view is commonly held in both physics and psychology (Aerts, 2009; Aerts & Sozzo, 2011; Bruza, Kitto, Nelson, & McEvoy, 2009; Busemeyer & Bruza, 2012; Filipp & Svozil, 2005; Khrennikov, 2009b; Yearsley & Pothos, 2014). In particular, among those applying quantum probability to behavior and also treating quantum probability theory as an alternative to CPT, there are claims that Bell-type inequalities are violated in experiments involving combinations of concepts (Aerts & Sozzo, 2011; Busemeyer & Bruza, 2012) and memory (Bruza, Kitto, Nelson, & McEvoy, 2009).

We will not recapitulate all the arguments related to this issue, as they have been presented in many previous publications (Dzhafarov, Cervantes, & Kujala, 2017; Dzhafarov & Kujala, 2014a, 2014b, 2016a, 2017a, 2017b; Dzhafarov, Kujala, & Larsson, 2015). We will use just one familiar example. Let R_1, R_2, R_3, R_4 denote a set of binary (+1/-1) random variables with known distributions of (R_1, R_2) , (R_2, R_3) , (R_3, R_4) , and (R_4, R_1) . The necessary and sufficient condition for the existence of such a quadruple of random variables is given by the CHSH/Fine inequality (Bell, 1964; Clauser, Horne, Shimony, & Holt, 1969; Fine, 1982):

$$\max_{j=1,\dots,4} \left| \sum_{i=1}^{4} \left\langle R_i R_{i\oplus 1} \right\rangle - 2 \left\langle R_j R_{j\oplus 1} \right\rangle \right| \le 2, \tag{1}$$

where $\oplus 1$ is cyclic shift $1 \to 2 \to 3 \to 4 \to 1$, and $\langle \cdot \rangle$ is expectation. One can easily construct examples of distributions of $(R_i, R_{i\oplus 1})$ for which this inequality is violated, indicating that such R_1, R_2, R_3, R_4 do not exist (essentially by the same logic as in determining that there are no four numbers a, b, c, d with a = b, b = c, c = d, and d = a + 1).

The problem arises when we are being told that the existence of such R_1, R_2, R_3, R_4 is predicted by quantum theory and corroborated by experiment. If we believe this, violations of (1) should indeed mean that CPT is inadequate, if not internally contradictory. We should not, however, believe this. R_1, R_2, R_3, R_4 in (1) are random variables in the CPT sense; they are not within the language of quantum theory. To decide what classical random variables should describe outcomes of what quantum measurements, one needs to go outside quantum theory. The general rule is that a random variable is identified by what is being measured and how it is being measured. The latter includes all conditions under which the measurement is made, in particular, all other measurements performed together with the given one. In our example, the measurements are indicated by star symbols in the following matrix:

$\overline{}$					١.
*	*			c_1	
	*	*		c_2	
		*	*	c_3	
*			*	c_4	
q_1	q_2	q_3	q_4		

The row labels c_1, \ldots, c_4 are called contexts, and here they are defined by which two quantities are being measured together: in c_1 it is q_1 and q_2 , in c_2 it is q_2 and q_3 , etc. In behavioral science the quantities q_1, \ldots, q_4 can be, e.g., four Yes-No questions posed to a large number of people divided into four groups: in the group c_1 each person is asked q_1 and q_2 , in the group c_2 each person is asked q_2 and q_3 , etc. In quantum mechanics the matrix above could describe the well-known EPR/Bell paradigm with two entangled spinhalf particles: q_1 and q_3 correspond to the two axes along which Alice measures spins in her particle, while q_2 and q_4 correspond to the two axes analogously used by Bob in his particle.

Let us use the notation R_i^j for the outcome of a measurement of q_i in context c_j :

R_1^1	R_2^1			c_1	
	R_2^2	R_3^2		c_2	
		R_3^3	R_4^3	c_3	
R_1^4			R_4^4	c_4	
q_1	q_2	q_3	q_4		

Since the values of R_i^i and $R_{i\oplus 1}^i$ are empirically paired (two responses given by the same person, or the measurements by Bob and Alice made simultaneously), the random variables in each row of the matrix are jointly distributed. This is not true for measurements made in different contexts: their joint distribution is undefined, and we call them stochastically unrelated (not to be confused with being stochastically independent, which is a special case of being jointly distributed). In particular, R_i^j and $R_i^{j'}$ measuring the same property q_i in two different contexts are stochastically unrelated. In the Kolmogorovian language, R_i^j and $R_i^{j'}$ are defined on different domain probability spaces. It is therefore impossible to say that $R_i^j = R_i^{j'}$, because equality is a special case of joint distribution.

It is clear now that CPT imposes no constraints whatever on the row-wise joint distributions. The CHSH/Fine inequality (1) cannot be derived for this matrix of contextually labelled random variables. However, it can be derived as a solution for the following problem: find necessary and sufficient conditions for the existence of a jointly distributed quadruple of random variables (S_1, S_2, S_3, S_4) such that

$$(S_i, S_{i \oplus 1})$$
 has the same distribution as $(R_i^i, R_{i \oplus 1}^i)$, (2)

for $i=1,\ldots,4$. Such a vector (S_1,S_2,S_3,S_4) is called a "reduced coupling" of the stochastically unrelated pairs $(R_i^i,R_{i\oplus 1}^i)$ (Dzhafarov & Kujala, 2016b).

The reduced coupling (S_1, S_2, S_3, S_4) is merely a shortcut for describing a special case of what we call a C-coupling

(Dzhafarov, Cervantes, & Kujala, 2017; Dzhafarov & Kujala, 2016a, 2017a, 2017b). C is some property of a pair of random variables, and a C-coupling of the pairs $\left(R_i^i, R_{i \oplus 1}^i\right)$ in our example is a jointly distributed octuple of random variables $\left(S_1^1, S_2^1, S_2^2, S_3^2, S_3^3, S_4^3, S_4^4, S_1^4\right)$ such that, for $i=1,\ldots,4$,

 $(S_i^i, S_{i \oplus 1}^i)$ has the same distribution as $(R_i^i, R_{i \oplus 1}^i)$, (3) and, in addition,

$$(S_{i\oplus 1}^i, S_{i\oplus 1}^{i\oplus 1})$$
 satisfies property C. (4)

The reduced coupling is the one defined by C with the meaning "are equal with probability 1" (applied to pairs of random variables). More generally, in the Contextuality-by-Default theory, C is chosen to mean "are equal with maximal possible probability." For this choice of C, the criterion for the existence of a C-coupling is

$$\max_{j=1,\dots,4} \left| \sum_{i=1}^{4} \left\langle R_{i}^{i} R_{i\oplus 1}^{i} \right\rangle - 2 \left\langle R_{j}^{j} R_{j\oplus 1}^{j} \right\rangle \right|$$

$$\leq 2 + \sum_{i=1}^{4} \left| \left\langle R_{i\oplus 1}^{i} \right\rangle - \left\langle R_{i\oplus 1}^{i\oplus 1} \right\rangle \right|,$$

$$(5)$$

a useful generalization of CHSH/Fine inequality (1) (Dzhafarov & Kujala, 2016a; Dzhafarov, Kujala, & Larsson, 2015; Kujala & Dzhafarov, 2016; Kujala, Dzhafarov, & Larsson, 2015). A system of random variables for which a C-coupling exists (does not exist) is called C-noncontextual (respectively, C-contextual).¹

That context is part of the identity of a random variable is the departure point for the Contextuality-by-Default theory, the term "identity" being understood in the Kolmogorovian sense, as the measurable function from a domain probability space to a codomain measurable space (for detailed explanations, see Dzhafarov & Kujala, 2016a, 2017a). One advantage provided by this contextual identification is that it allows for the possibility that random variables measuring the same property in different contexts, such as R_2^1 and R_2^2 in our example, are differently distributed. This can happen, e.g., if one of the two questions posed to a person influences her response to the other question, or if Alice can signal to Bob and thereby change his recordings. With noncontextual labeling, such as R_1, R_2, R_3, R_4 in the opening formulation, to express the same fact one would have to say that R_2 is differently distributed depending on whether "it" is recorded together with R_1 or R_3 . This is at best an abuse of language, if not outright nonsensical, as the distribution of R_2 is part of its identity.

2. ON STATEMENT 2

"In classical probability the joint occurrence of two events is commutative, but we know from

To avoid terminological confusion, in the Contextuality-by-Default theory each random variable is contextually labeled ("by default"), but a system of contextually labelled random variables can be Cnoncontextual or C-contextual.

quantum mechanics and from behavioral experiments that the order of two events generally matters for their joint probability."

Thus, we read in Trueblood and Busemeyer (2011) that "the classical probability model has difficulty accounting for order effects because the commutative property holds" (p. 1527). And in Wang and Busemeyer (2015): "Classical probability theory has difficulty explaining order effects because events are represented as sets and are commutative, so the joint probability of events A and B is the same for the order of 'A and B' and the order of 'B and A" (p. 2). Quotes like these are numerous, but it should be noted that Busemeyer and colleagues carefully qualify their criticism of CPT. They acknowledge that models based on CPT can be formulated for such empirical phenomena as order effects, but their presentation implies that these CPT-based models have to be contrived. According to these authors, the only way CPT can handle order effects is by using the Ballentine (1986) type conditionalization: order of events ("B follows A" and "A follows B") is considered a random event conditioning probabilities of responses. We too consider this construction awkward (Dzhafarov & Kujala, 2014b; see also Section 3), but it is not the only one within the framework of CPT: the Contextuality-by-Default approach provides another way, one that is both simple and universally applicable.

Let us precede our discussion by pointing out that CPT would indeed be a singularly helpless exercise if it lacked natural ways to depict the difference between an ordered pairs of observations (a, b) and an unordered two-element set $\{a,b\}$. The difference between the two is obvious on the basic set-theoretic level: an ordered pair (a, b) is an abbreviation for the set $\{\{a,1\},\{b,2\}\}\$, or $\{a,\{a,b\}\}\$, because of which (a,b) and (b,a) are different sets, unless a=b, and $(a, a) = \{\{a, 1\}, \{a, 2\}\}\$ is different from $\{a, a\} = \{a\}$. Moreover, since an ordered pair is merely a simple case of a process (indexed set), the logic of Statement 2 implies that CPT should resort to contrived constructions when dealing with random processes that are not exchangeable. A statement from Bruza, Wang, and Busemeyer (2015) may help in recognizing that the order effects are a non-issue for CPT. The statement is that in CPT "the intersection of events is always defined and events always commute, even if the events are distinguished by time (e.g., 'A at time 1' and 'B at time 2' is equivalent to 'B at time 2' and 'A at time 1')" (p. 387). For ordered events it is precisely this commutativity, of 'A at time 1' and 'B at time 2', that holds in CPT (and classical logic). In this form it is unchallengeable and cannot lead to any problems. An issue is created when one compares 'A at time 1' and 'B at time 2' to 'B at time 1' and 'A at time 2', two conjunctions that are two different events that need not have the same probability in

However, the approach offered by the Contextuality-by-Default theory does not consist in labeling events. Rather, it uses a more versatile labeling of random variables (although the two are essentially equivalent in the simple case of two consecutive events). To understand the logic of the

approach, consider the probabilistic identity of responses R_q to some question q. Its domain probability space can be thought of as a set X of potential responders to the question q, with some probability measure μ imposed on its power set (treated as sigma-algebra). Let the possible values of R_q be Yes/No. Its distribution then is defined by

$$\Pr\left[R_q = \text{Yes}\right] = \mu\left(\left\{x \in X : x \text{ responds Yes to } q\right\}\right). \quad (6)$$

By construction, q is part of the identity of R_q , so if q is replaced with another question q', the random variable R_q will be replaced with another random variable $R_{q'}$. Probability theory allows this new random variable to have another distribution, but, of course, being an abstract mathematical theory, it does not predict what the distributions of R_q and $R_{q'}$ can be: such a prediction is up to an empirical theory dealing with people's substantive knowledge of questions and answers.

Consider now two questions that have identical formulation but are asked in different tones of voice or with different noise or images in the background; or two questions that have the same content but differ in how they are formulated (e.g., "Is it 11 am now?" versus "Is 11 am the correct time at this moment?"). The usual experimental design, if one is interested in such differences, would be to partition X into two subsets X^1 and X^2 , asking the question q in one form of the members of X^1 and in another form of the members of X^2 . From the point of view of abstract probability theory, whatever the difference between the two questions substantively, formally the responses to them are two different random variables defined on two different domain probability spaces. They are, therefore, stochastically unrelated. One can choose (based on one's substantive, nonmathematical understanding of questions and answers) to consider the differences in formulations or in the tone of voice to be part of the questions themselves (in which case one will deal with random variables denoted R_q and $R_{q'}$) or to formalize the differences as different contexts in which one and the same question is asked (in which case one will present the random variables as R_q^c and $R_q^{c'}$). The two representations are interchangeable, but the latter one is preferable because, in accordance with the principles of the Contextuality-by-Default theory, it encodes the stochastic unrelatedness of R_q^c and $R_q^{c'}$ in the very notation.²

Using different orders of two questions has precisely the same logical status as differences in the tone of voice or background noise: it creates two pairs of jointly distributed random variables that are stochastically unrelated to each other. The set X is partitioned into two subsets X^{AB} and X^{BA} , corresponding to the two orders, (q_A, q_B) and (q_B, q_A) . The random variables defined on these subsets

 $^{^2}$ A combined notation, such as R_q^c and $R_{q^\prime}^{c\prime}$, is possible too, as discussed by Dzhafarov and Kujala (2015), but it is less interesting for subsequent contextuality analysis. As a general principle, it is possible but counterproductive to include contexts as part of contents of random variables: strictly separating the two is essential for any contextuality analysis.

and corresponding to a given question, say q_A , can have different distributions. The latter is exceedingly obvious if one uses specially chosen questions. Consider, e.g., q_A ="Is this the first question I am asking?" and q_B ="Is this the second question I am asking?", asked in two different orders.

By analogy with two forms of the same question, one can now proceed in several different ways, but the one most informative for contextuality analysis is as follows. We define a jointly distributed pair (R_A^{AB}, R_B^{AB}) with

$$\Pr\left[R_A^{AB} = \text{Yes}\right] = \mu\left(X_A^{AB}\right),$$

$$\Pr\left[R_R^{AB} = \text{Yes}\right] = \mu\left(X_R^{AB}\right),$$
(7)

$$\Pr\left[R_A^{AB} = \text{Yes \& } R_B^{AB} = \text{Yes}\right] = \mu\left(X_A^{AB} \cap X_B^{AB}\right),$$

where

$$X_A^{AB} = \left\{ x \in X^{AB} : x \text{ responds Yes to } q_A \right\},$$

$$(8)$$

$$X_{B}^{AB}=\left\{ x\in X^{AB}:x\text{ responds Yes to }q_{B}\right\} .$$

The joint distribution for (R_A^{BA}, R_B^{BA}) , stochastically unrelated to the previous pair, is defined similarly, and can be arbitrarily different from (7).

Using the Contextuality-by-Default representation, the system of the random variables just defined is

$$\begin{vmatrix} R_A^{AB} & R_B^{AB} & c_{AB} = (q_A, q_B) \\ R_A^{BA} & R_B^{BA} & c_{BA} = (q_B, q_A) \\ q_A & q_B \end{vmatrix} .$$

We can now choose some property C for pairs of random variables, as explained in the previous section, and ask whether the system above has a C-coupling (or, in the terminology of Contextuality-by-Default, whether it is C-noncontextual). With C chosen to mean "are equal with maximal possible probability," such a C-coupling exists if and only if (Dzhafarov & Kujala, 2016a; Dzhafarov, Zhang, & Kujala, 2015)

$$\begin{aligned} \left| \left\langle R_{A}^{AB} R_{B}^{AB} \right\rangle - \left\langle R_{A}^{BA} R_{B}^{BA} \right\rangle \right| \\ &\leq \left| \left\langle R_{A}^{AB} \right\rangle - \left\langle R_{A}^{BA} \right\rangle \right| + \left| \left\langle R_{B}^{AB} \right\rangle - \left\langle R_{B}^{BA} \right\rangle \right|. \end{aligned} \tag{9}$$

Here, Yes and No responses have been encoded as +1 and -1, respectively. The remarkable QQ equality discovered by Wang and Busemeyer (2013) is equivalent to saying that the left-hand side expression in (9) is zero, from which it follows that according to this law this system of random variables is noncontextual.³ See Dzhafarov, Kujala, Cer-

vantes, Zhang, and Jones (2016) and Dzhafarov, Zhang, and Kujala (2015) for a detailed discussion.

3. ON STATEMENT 3

"Classical probability is additive (equivalently, obeys the law of total probability), but we know from quantum mechanics and from behavioral experiments that this additivity (the law of total probability) can be violated."

Additivity, expressed in the language of random variables, is that if A and B are disjoint events in the codomain space of a random variable R, then

$$\Pr\left[R \in A \cup B\right] = \Pr\left[R \in A\right] + \Pr\left[R \in B\right]. \tag{10}$$

This principle is sometimes analyzed in an equivalent form, referred to as the "law of total probability": as a consequence of the additivity above and the set-theoretic distributivity, for any C in the codomain space of R,

$$\Pr[R \in C \cap (A \cup B)]$$

$$= \Pr[R \in C \cap A] + \Pr[R \in C \cap B].$$
(11)

Set-theoretic distributivity being an integral part of CPT, and the two equalities above understood as belonging to CPT, it is logically impossible to claim that one of them can be violated without stating the same for the other. This should be kept in mind when encountering statements about violations of the "classical law" of total probability. To give examples: "It was shown that FTP [the formula of total probability] (and hence classical probability theory) is violated in some experiments on recognition of ambiguous pictures" (Khrennikov, 2010, p. 90), and "One can find evidence of violation of laws of classical probability theory, e.g., in violation of the law of total probability" (Khrennikov & Basieva, 2014, p. 105).

We will discuss here the basic form (10) only.

The claim of violations of this law in quantum mechanics comes from the double-slit experiment. We consider it in the following version: a source of particles emits them into a barrier with two slits (left and right, each of which can be closed or open), and a detector of the particles occupies a small area behind this barrier. One considers the probability with which an emitted particle reaches the detector, and discovers that this probability, when both slits are open, is not equal to (depending on the detector's location, can be greater or smaller than) the sum of these probabilities recorded with only the left slit open and with only the right slit open. Richard Feynman is often quoted as saying that this is "a phenomenon which is impossible, absolutely impossible, to explain in any classical way" (Feynman, Leighton, & Sands, 1975, Section 37-1). The words "in any classical way" in this quote are commonly interpreted as "by means of CPT." This interpretation may be correct, but it is also possible that Feynman meant that

³ Wang and Busemeyer (2013) are right about CPT (they call it "Bayesian") not being able to predict the QQ equality, although they seem to consider this a deficiency rather than a hallmark of any abstract mathematical theory. Any prediction derived in a special-purpose theory, such as the quantum formalism used by Wang and Busemeyer to derive the QQ equality, can be (and always is, eventually) fully expressed in the language of CPT, making it possible to relate the prediction to empirical data.

this phenomenon cannot be explained by means of classical mechanics, and that he viewed quantum probabilities as a special-purpose theory for computing probabilities in a specific physical situation. The second quote from Feynman (1951) given at the end of our introductory section seems to agree with this interpretation.

Whatever the case with Feynman, Ballentine (1986) presents a systematic analysis of the double-slit experiment in terms of CPT, and argues that the two are perfectly compatible if one treats the probabilities in (10) as conditional ones, conditioned on three different events. Translating this into the language of random variables, Ballentine's solution is to rewrite (10) as

$$\Pr\left[R \in A \cup B \mid Q = c_{\circ \circ}\right]$$

$$\stackrel{generally}{\neq} \Pr\left[R \in A \mid Q = c_{\circ \times}\right] + \Pr\left[R \in B \mid Q = c_{\times \circ}\right],\tag{12}$$

where Q is a random variable indicating which of the slits is open and which is closed: $c_{\circ\circ}$ means that both are open, $c_{\circ\times}$ means that only the left one is open, and $c_{\times\circ}$ means the opposite. Clearly, the three conditioning values $c_{\circ\circ}, c_{\circ\times}, c_{\times\circ}$ are distinct and mutually exclusive (as with any distinct values of any random variable), whence no equality in (12) should generally be expected.⁴

Dzhafarov and Kujala (2014b) call this approach "conditionalization," in relation to Avis, Fischer, Hilbert, and Khrennikov (2009) where it was used systematically (see also Khrennikov, 2006, 2015b). It is true that if R and Q are jointly distributed, then R conditioned on some value of Q and R conditioned on another value of Q are two random variables that possess no joint distributions, i.e., are stochastically unrelated. This means that conditionalization is a special case of contextual labeling, in fact an instructive case for introducing the notion of stochastic unrelatedness (Dzhafarov & Kujala, 2014b, 2016b). However, the choice among conditions $c_{\circ\circ}, c_{\circ\times}, c_{\times\circ}$ need not be random. One can conduct an experiment with both slits open for a year, then for another year with the left slit closed,

and so on. This should not change anything in the analysis of the double-slit experiment. As we mentioned in the introductory section, Khrennikov (2009a) pointed out the difference between contextual and conditional probabilities in presenting his general Växjö contextual model.⁵

The analysis of the double-slit experiment within the framework of the Contextuality-by-Default theory begins with identifying the random variables in play, their contexts and the properties they measure. The contexts are the same as in Ballentine's analysis, $c_{\circ\circ}, c_{\circ\times}, c_{\times\circ}$, but for completeness we add $c_{\times\times}$ (both slits closed). The measured properties are identified by which slit one considers (left or right) and by its condition (open or closed): q_{\circ} . (left slit, open), $q_{\circ\circ}$ (right slit, open), and similarly for the closed slits, q_{\times}, q_{\times} . This creates eight random variables that we can arrange as follows:

R_{\circ}°	$R^{\circ\circ}_{\cdot\circ}$			$c_{\circ\circ}$	
	$R^{\times \circ}_{\cdot \circ}$	R_{\times}°		$c_{\times \circ}$	
		R_{\times}^{\times}	$R_{\cdot \times}^{\times \times}$	$c_{\times \times}$,
$R_{\circ}^{\circ \times}$			$R^{\circ \times}_{\cdot \times}$	$c_{\circ \times}$	
q_{\circ} .	$q_{\cdot \circ}$	q_{\times} .	$q_{\cdot imes}$		

where R_q^c can be interpreted as answering the question "Has the particle just emitted reached the detector through the slit q under condition c?" The possible values of R_q^c are Yes and No. Thus, $R_{\times}^{\times \circ}$ =Yes means that a particle reached the detector having passed through the closed slit on the left when the right slit is open. Physics (not probability theory) tells us that the probability of this happening is zero.

Our arrangement of the random variables shows that, surprisingly, the system they comprise is formally a cyclic system of rank 4 (Dzhafarov & Kujala, 2016a; Dzhafarov, Kujala, & Larsson, 2015; Kujala, Dzhafarov, & Larsson, 2015). It is the same system as the one in the simplest EPR/Bell "Alice-Bob" paradigm, described in our analysis of Statement 1. If one uses one's knowledge that no particle can reach the detector through a closed slit, then the joint distributions of all context-sharing pairs of random variables (the rows of the matrix above) are defined by the following joint and marginal probabilities:

$$\begin{array}{c|c|c} c_{\circ\times} & R_{\cdot\times}^{\circ\times} = \mathrm{Yes} & c_{\times\times} & R_{\cdot\times}^{\times\times} = \mathrm{Yes} \\ \hline R_{\circ\cdot}^{\circ\times} = \mathrm{Yes} & 0 & p \\ \hline 0 & & 0 \\ \end{array} , \begin{array}{c|c|c} c_{\times\times} & R_{\cdot\times}^{\times\times} = \mathrm{Yes} \\ \hline R_{\times\cdot}^{\times\times} = \mathrm{Yes} & 0 & 0 \\ \hline 0 & & 0 \\ \end{array} .$$

$$\begin{array}{c|c} c_{\times \circ} & R_{\circ}^{\times \circ} = \mathrm{Yes} \\ \hline R_{\times}^{\times \circ} = \mathrm{Yes} & 0 & 0 \\ \hline q & & r' + q' \\ \end{array},$$

⁴ In a personal communication (April 2018), Jerome Busemeyer explained to us that when he and his colleagues speak of violations of the total probability law they do not mean that (10) or (11) fail to hold. Rather they mean that the probability of an event can be different depending on what other events it is recorded together with. This can be interpreted as a position very close to Contextualityby-Default, or to Ballentine's conceptualization (12), or even as the possibility that, say, C when one measures the probability of $R \in C \cap A$ may be a different event from C when one measures the probability of $R \in C \cap (A \cup B)$ (the latter case being formulated as measuring the probability of "C alone" if B = notA). While welcoming this clarification, one should note that its implication is that the law of total probability as a formula of CPT is not violated. Ballentine's inequality (12) is a correct CPT formula, and something like $\Pr[R \in C_1 \cap (A \cup B)] = \Pr[R \in C_2 \cap A] + \Pr[R \in C_3 \cap B]$ is not a correct CPT formula. (The latter, incidentally, is a good example of why labeling of events, rather than of random variables, is not a good solution: shall one, in addition to C, also differently label A and B on the right and on the left of the formula?)

⁵ In some of his later work, however, Khrennikov seems to have abandoned this distinction and adopted a rigorous version of Ballentine's view (Avis, Fischer, Hilbert, & Khrennikov, 2009; Khrennikov, 2015a, 2015b; see Dzhafarov & Kujala, 2014b, for a critical discussion of this position).

where p, q, p', q', r' are some probabilities. The physical interpretation of the joint distribution for $c_{\circ\circ}$ compared to that for, say, $c_{\circ\times}$ is that, somehow, the way particles reach the detector having passed through the open left slit may be different depending on whether the right slit is open or closed. A physicist may tell us that this is because of the particle-wave duality and wave interference, but this is irrelevant for the probabilistic analysis.

It is interesting to see whether the system just described is C-noncontextual (has a C-coupling) with C="are equal with maximal possible probability." The application of the general criterion (5) for noncontextuality of such a system yields

$$((1-2p) + (1) + (1-2q) + (1-2p'-2q'))$$

$$-2\min((1-2p), (1), (1-2q), (1-2p'-2q'))$$

$$\leq 2+2|p-p'-r'|+|-1+1|$$

$$+2|q-q'-r'|+|-1+1|,$$
(13)

where we have assumed that the detector is so small that the probabilities 1-2p, 1-2q, 1-2p'-2q' are all positive. By simple algebra one can show that this inequality is always satisfied, that is, the double-slit system is C-noncontextual.

Comparing again $c_{\circ\circ}$ with $c_{\circ\times}$ (or $c_{\times\circ}$), the noncontextuality just established means, within the framework of Contextuality-by-Default, that the influence exerted by the state of a slit (open or closed) upon how the particles reach the detector having passed through the other slit is of a "direct cross-influence" nature, with no contextuality proper (Dzhafarov, 2017; Dzhafarov, Cervantes, & Kujala, 2017; Dzhafarov & Kujala, 2017a). For a detailed contextuality analysis of the double-slit experiment see Dzhafarov and Kujala (2018), where it is also shown that a system with more than two slits may very well be contextual.

4. CONCLUDING REMARKS

In this concluding section we will briefly address four commonly raised concerns about the contextual notation and the principle that random variables in different contexts are different (and stochastically unrelated).

Question: In empirical situations where the contexts are known not to influence a measurement directly (like in the EPR/Bell Alice-Bob paradigm with spacelike separation of the measurements), what "causes" the random variable representing this measurement to change its identity?

Answer: The identity of a random variable is determined by its own distribution and also by the joint distribution of this random variable with all other random

variables in the same context. Therefore, any change in these other variables "automatically" changes its identity. Here is a simple analogy. A person P is in a room with other people. P has some characteristics, such as "she is kind," or "she is tall." It is possible that she is the tallest person in the room, in which case she is also characterized by this fact. The statement "she is the tallest person in the room" therefore describes a property of P, part of her identity in addition to her being kind and tall. If one of the other people leaves, and someone enters who is taller than P, she "automatically" changes her identity, as she ceases to be the tallest person in the room. This change in Poccurs even if she is not aware of the change in the room, or the room is so large that there are no physical means for her to notice this. As acknowledged in Dzhafarov and Kujala (2014a), an application of this general argument to quantum phenomena may be regarded as paralleling an argument made by Bohr (1935) in reply to the EPR paper (Einstein, Podolsky, & Rosen, 1935).

Question: If every condition recorded together with a random variable can be considered part of its context, does this not mean that any two realizations of the same random variable are in fact two different random variables, stochastically unrelated to each other?

Answer: If the realizations are separately indexed, e.g., by the ordinal position in a sequence of trials, each of them indeed must be viewed as a single realization of a unique random variable. There is, however, a choice of the point of view for subsequent analysis. One can view these unique random variables as ones with different measured properties (trial numbers) within a single context (sequence of trial numbers). Conversely, one can view them as random variables measuring the same thing (e.g., they all measure the response of a person to a flash) but in different contexts (trial number). We implicitly adopt the second point of view when we speak of the sequence as one of different realizations of "the same" random variable. See Dzhafarov and Kujala (2015, 2016a) for detailed discussions. In the case considered, the choice of one of the two points of view makes no difference. If the realizations are treated as context-sharing, the random variables are jointly distributed, but this joint distribution is manifested in a single realization only. We need additional assumptions to reconstruct it, such as stochastic independence, ergodicity, or martingale property. If the realizations are treated as measuring the same property in different contexts, they are pairwise stochastically unrelated, and we need to couple them. The choice of a coupling here amounts to adopting the same additional assumptions.

Question: Is the contextual labeling with stochastic unrelatedness really classical, in the Kolmogorovian sense? Answer: It is a matter of definition and understanding of history. In some publications one of us and Janne Kujala called our approach a "qualified" Kolmogorovian theory (Dzhafarov & Kujala, 2014c), and it can be presented in a way that sets it aside from a standard account of CPT (as in Dzhafarov & Kujala, 2016a). However, we prefer to speak

of Contextuality-by-Default as part of the Kolmogorovian probability theory, with a greater emphasis on multiple freely introducible domain probability spaces, stochastically unrelated random variables defined on these spaces, and their couplings understood as placing their copies on the same domain space. This preference is based on our disbelief that Kolmogorov himself and the brilliant probabilists working in his language could have overlooked the obvious fact that there cannot exist a joint distribution of all imaginable random variables. In his celebrated little book, Kolmogorov (1956, §2 of Chapter 1) discusses empirical applications of his mathematical theory, and in doing so confines his consideration to a single experiment (corresponding, in our language, to a single context). He may have erroneously thought this was the only realistic or interesting application. This reading of Kolmogorov is also advocated in Khrennikov (2009b).

Question: If, however, one posits that in any application of CPT all random variables involved are defined on a single domain probability space, would not then the claim of the inadequacy of CPT thus understood be justified?

Answer: The issue of the existence of a joint distribution for all random variables involved in a given application is not as critical as the issue of what random variables are involved. Statements 1, 2, and 3 considered above are based first and foremost on misidentifying the random variables in play. Thus, the correct system of random variables representing the question order experiment is

$$\mathsf{R} = \begin{bmatrix} R_A^{AB} & R_B^{AB} & c_{AB} = (q_A, q_B) \\ R_A^{BA} & R_B^{BA} & c_{BA} = (q_B, q_A) \\ \hline q_A & q_B \end{bmatrix}.$$

It is simply unjustifiable to posit a priori that it can be replaced with

$$\mathsf{R}' = \begin{bmatrix} R_A & R_B & c_{AB} = (q_A, q_B) \\ R_A & R_B & c_{BA} = (q_B, q_A) \\ q_A & q_B \end{bmatrix},$$

a system in which random variables do not change with context: even if one ignores the logic of Contextuality-by-Default, there is no rationale for assuming that contexts are irrelevant, because in this particular example one even knows that the distributions of R_A^{AB} and R_A^{BA} are different (which is the very "question order effect" that makes this paradigm interesting). The situation here is no different from someone deciding to replace R with

$$\mathsf{R}'' = \begin{bmatrix} R^{AB} & R^{AB} & c_{AB} = (q_A, q_B) \\ R^{BA} & R^{BA} & c_{BA} = (q_B, q_A) \\ q_A & q_B \end{bmatrix},$$

a system in which random variables do not change with content.

This reasoning applies even if one views the four random variables in the original system R as having an unknown (and unknowable) joint distribution. This amounts

to informally identifying the system with one of its possible couplings, and the construction of a C-coupling then can be presented as determining if this "true" joint distribution could possibly be satisfying C. One can check that our analysis of the question order effect in Section 2 would hold with no serious modifications if one adopted this language (and similarly for the systems considered in Sections 1 and 3).

There are, of course, good reasons not to use this language, except as an informal version of the rigorous language of the Contextuality-by Default theory (perhaps for the sake of conceptual or notational simplicity). The assumption that any two random variables are jointly distributed is mathematically untenable. It is untenable because, due to the transitivity of the relation of being defined on the same domain probability space, it implies the erroneous notion that there is a joint distribution for all imaginable random variables (for reasons why this notion is wrong, see Dzhafarov & Kujala, 2014a, 2014b, 2017a). Within the framework of CPT stochastically unrelated random variables must exist, making it unjustifiable to assume without critical examination that all random variables in a given set have a joint distribution.

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